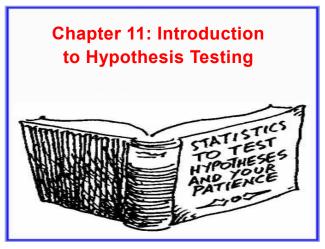
2



#### Introduction

- In this chapter we present the second general procedure of making inferences about a population – hypothesis testing.
- The purpose of hypothesis testing is to determine whether we have enough statistical evidence to conclude that a belief or hypothesis about a parameter is supported by the data.
- This chapter represents a critical contribution to your development as a statistics practitioner because it lays the foundation for the rest of the properties.
- First, we introduce the concepts of hypothesis testing, and then we develop a
  method to test a hypothesis about a population mean when the population
  standard deviation is known.

2

#### **Concepts of Hypothesis Testing**

The critical concepts in **hypothesis testing** may be summarized as follows:

- 1. There are two hypotheses
  - The  ${\bf null\ hypothesis}\ (H_{\rm o}$  pronounced " $H\ nought$ ")
  - The **alternative** (or research) **hypothesis** (*H*<sub>1</sub>).
- 2. The test begins with the assumption that  $H_{\rm o}$  is true.
- 3. The goal of the test is to determine whether there is enough evidence to infer that  ${\cal H}_i$  is true.
- 4. There are two possible decisions
  - \* Reject  $H_o$ . Conclude that there is enough evidence to support  $H_r$ .
  - Do not reject  $H_o$ . Conclude that there is not enough evidence to support  $H_i$ .

Nonstatistical Application of Hypothesis Testing

In a criminal trial, a defendant stands accused of a crime in front of a jury.

It can be said that the jury conducts a test of hypotheses based on the evidence presented by both the prosecution and the defense, that is:

- $H_0$ : The defendant is innocent.
- H<sub>1</sub>: The defendant is guilty.

The jury can only make two decisions (i.e., render two verdicts):

- Reject H<sub>0</sub>. There is enough evidence to conclude that the defendant is guilty: a conviction.
- Do not reject H<sub>0</sub>. There is not enough evidence to conclude that the defendant is guilty: an acquittal.

Under no circumstances, the decision that  $H_0$  is true, and the defendant is innocent, is ever made.

4

3

### **Error Types**

Two possible errors can be made in a test:

 A Type I error occurs when we reject a true null hypothesis.

 $P(\text{Type I error}) = \alpha$ , also called the **significance level**.

 A Type II error occurs when we do not reject a false null hypothesis.

 $P(\text{Type II error}) = \beta$ 

The error probabilities  $\alpha$  and  $\beta$  are inversely related, and any attempt to reduce one will increase the other. The table on the next slide summarizes the terminology and the concepts.

5

6

#### **The Blackstone Ratio**

- Justice system (in UK and Canada) is arranged by placing the burden of proof on the prosecution.
- The judges instruct the jury to find the defendant guilty only if there is "evidence beyond a reasonable doubt", and in the absence of enough evidence, the jury must acquit.
- It follows that the probability of acquitting guilty people (Type II error) is greater than the probability of convicting innocent people (Type I error.)
- Sir William Blackstone (1723–1780), an English legal scholar, famously quipped: "Better for 10 guilty persons escape than that one innocent suffers." This is known as the Blackstone Ratio.

The Blackstone Ratio:

 $\frac{P(\text{Type I error})}{P(\text{Type II error})} = \frac{\alpha}{\beta} = \frac{1}{10}$ 

Benjamin Franklin (1706–1790) thought that a 1 to 10 ratio was not enough when he stated: "it is better 100 guilty persons should escape than that one innocent person should suffer, is a maxim that has been long and generally approved."

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#### **Statistical Application of Hypothesis Testing**

Recollect the example of the operations manager at Doll's Computers wanted to estimate the mean demand during lead time. Suppose that the manager wants to know instead whether the mean demand is *different from* 350.

Now the question becomes: "Is there enough evidence to conclude that  $\mu$  is *not equal to* 350?"

The hypotheses are set up to reflect a manager's decision problem wherein the null hypothesis represents the  $status\ quo$ , whereas the  $burden\ of\ proof$  is placed on the alternative hypothesis.

Thus, we write:  $H_o$ :  $\mu = 350$ 

 $H_1$ :  $\mu \neq 350$ 

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	Population Condition	
Conclusion	$H_0$ True $(\mu \le 12)$	H <sub>0</sub> False (μ > 12)
Accept $H_0$ (Conclude $\mu \le 12$ )	Correct Decision	Type II Error
Reject $H_0$ (Conclude $\mu > 12$ )	Type I Error	Correct Decision

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#### **Statistical Application of Hypothesis Testing**

And what if the manager wants to know whether the mean demand has increased from 350?

The new question is now: "Is there enough evidence to conclude that  $\mu$  is greater than 350?"

We can write the hypotheses as:

 $H_o$ :  $\mu = 350$ 

 $H_1$ :  $\mu > 350$ 

What if the manager wants to know whether the mean demand has *decreased* from 350?

We can phrase the question as: "Is there enough evidence to conclude that  $\mu$  is less than 350?"

And write the hypotheses as:

 $H_o$ :  $\mu = 350$ 

 $H_1$ :  $\mu$  < 350

.

10

#### A Statistical Application of Hypothesis Testing

Finally, suppose that the manager does not know the actual mean demand, but the current inventory policy is based on the assumption that the mean is *less or equal to 350*. He wants to know whether the mean demand has *increased from 350*?

And write the hypotheses as:

 $H_0: \mu \le 350$ 

 $H_1$ :  $\mu > 350$ 

We always assume the actual form:

 $H_o$ :  $\mu = 350$ 

 $H_1$ :  $\mu > 350$ 

**Reason:** If the alternative hypothesis is true when  $\mu = 350$  then it is true when  $\mu \le 350$ 

#### **The Test Statistic**

The **test statistic** is based on the best estimator of the population parameter, and it is the criterion used to make our decision about the hypotheses.

In the Doll example, the test statistic is the mean of a random sample drawn from the population.

We use the test statistic to make one of the following two decisions:

- If the test statistic's value is inconsistent with the hypothesized mean of 350, we reject H<sub>0</sub> and conclude that there is sufficient evidence that H<sub>1</sub> is true.
- If the test statistic's value is close to the hypothesized mean of 350, we do not reject H<sub>o</sub> and conclude that there is no sufficient evidence that H<sub>i</sub> is true.

In the next section, we show how to use the test statistic's sampling distribution to define "sufficient evidence."

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## Estimating the Population Mean when the Population Standard Deviation is Known

Example - Department Store's New Billing System

After a thorough financial analysis, the manager of a department store has determined that a new billing system will be cost-effective only if the mean monthly account is more than \$170.

A random sample of 400 monthly accounts has a sample mean of \$178. The manager knows that the accounts are approximately normally distributed with a standard deviation of \$65.

Question: "can the manager conclude that the new system will be cost-effective?" The null hypothesis can be expressed as:

 $H_0$ :  $\mu$  ≤ 170

However, as discussed in Section 11-1, we actually test  $\mu$  = 170, so we set the hypotheses as:

 $H_o$ :  $\mu$  = 170 (system is not cost-effective: do not install new system)

 $H_1$ :  $\mu > 170$  (system is cost-effective: install new system)

12

11 12

#### **Defining the Rejection Region**

Example- Department Store's New Billing System

To conduct the test, we ask: "is a sample mean of 178 sufficiently greater than 170 to conclude that the population mean is greater than 170?"

To make a decision about the sample mean, we set up the rejection region: The rejection region is a range of values such that if the test statistic falls into that range, we decide to reject  $H_0$  in favor of  $H_1$ .

We define the rejection region such that:

 $\bar{x} > \bar{x}_L$ 

 $\mu=170$   $X_{i}$  Rejection region

Because a Type I error is defined as rejecting a true null hypothesis, we can write that:

 $P(\text{Type I error}) = P(\bar{x} > \bar{x}_L \mid H_0 \text{ is true}) = \alpha$ 

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#### Calculating the Test Statistic from the Rejection Region

Example - Department Store's New Billing System

From before, we know that  $\bar{X}$  is normally distributed. As a result, we can standardize  $\bar{x}$ :

$$P\left(\frac{\bar{x}-\mu}{\sigma/\sqrt{n}} > \frac{\bar{x}_L - \mu}{\sigma/\sqrt{n}}\right) = P\left(Z > \frac{\bar{x}_L - \mu}{\sigma/\sqrt{n}}\right) = \alpha$$

From Section 8.2, we defined  $z_{\alpha}$  such that:

$$P(Z > z_{\alpha}) = a$$

Because the two probabilities have the same distribution, they also have the same limit:

$$\frac{\bar{x}_L - \mu}{\sigma / \sqrt{n}} = z_\alpha$$

Solving for  $\bar{x}_L$  yields:

$$\bar{x}_L = \mu + z_\alpha \, \sigma / \sqrt{n}$$

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13 14

#### Calculating the Test Statistic from the Rejection Region

Example - Department Store's New Billing System

We can choose a significance level,  $\alpha=5\%$ , that is,  $z_{\alpha}=1.645$ . If now we plug in  $\mu=170$ ,  $\sigma=65$ , n=400, we obtain:

$$\bar{x}_L = 170 + 1.645 \left(65/\sqrt{400}\right) = 175.34$$

Therefore, the rejection region is:

 $\bar{x} > 175.34$ 

Because the test statistic (sample mean) is 178, it is in the rejection region.

Thus, we reject the null hypothesis, and conclude that there is sufficient evidence that the mean monthly account is greater than 170.

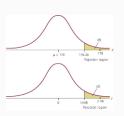
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#### Standardized Test Statistic

• One limitation of the previous method is that the rejection region is set up in terms of  $\bar{x}$ . It is easier to use the **standardized test statistic** instead:

$$z > \frac{\bar{x} - \mu}{\sigma / \sqrt{n}}$$

- Algebraically, the rejection region can be written as:  $z>z_{\alpha}$
- In our example we have:  $z>z_{\alpha}=z_{.05}=1.645$
- The standardized test statistic is:
- $z > \frac{\bar{x} \mu}{\sigma / \sqrt{n}} = \frac{178 170}{65 / \sqrt{400}} = 2.46$
- Because 2.46 is greater than 1.645, we reject the null hypothesis and conclude that there is enough evidence to infer that the mean monthly account is greater than \$170.
- For simplicity, we will refer to the standardized test statistic as the test statistic for the rest of this course.



Example - Department Store's

New Billing System

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#### p-Value

The main limitation of the rejection region method is that it produces a yes or no answer to the question of whether to reject  $H_o$  but offers no further information on the statistical significance of the test.

One of the implications of using the rejection region method is the impossibility of calculating the exact possibility of making an error (or the exact probability that a Type I error will occur).

A better approach that provides a measure of the am statistical evidence supporting  $H_i$  uses the p-value.

The **p-value** of a test is the probability of observing a was statistic at least as extreme as the one computed given that the null hypothesis is true.

In our example, the p-value is the probability of observing  $\bar{x} \ge 178$  when  $\mu = 170$  is:  $p\text{-value} = P(\bar{X} > 178) = P\left(\frac{\bar{X} - \mu}{\sigma/\sqrt{n}} > \frac{178 - 170}{65/\sqrt{400}}\right) = P(Z > 2.46) = 1 - P(Z < 2.46) = 1$ 

1 - .9931 = .0069

1'

#### Interpreting the p-Value

Example - Department Store's New Billing System

We just concluded that the probability of observing  $\bar{x} \ge 178$  when  $\mu=170$  is .0069, which is a very unlikely event.

Consequently, we have reason to reject  $H_o$  and support  $H_i$ . But that we cannot make a probability statement about a parameter that is not a random variable.

**Incorrect interpretation:** the p-value is the probability that the  $H_o$  is true. **Correct interpretation:** the p-value measures the amount of statistical evidence

that supports  $\hat{H}_{i}$ .

The table and the graph to the right show several values of  $\bar{x}$ , their z-statistics, and p-values.

Notice how the farther  $\bar{x}$  is from the hypothesized mean of 170, the smaller the p-value is, and the more the statistical evidence supports  $H_1$ .



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#### Summary of Suggested Guidelines for Interpreting *p*-Values

Less than .01

Overwhelming evidence to conclude  $H_a$  is true.

Between .01 and .05

Strong evidence to conclude  $H_a$  is true.

■ <u>Between .05 and .10</u>

Weak evidence to conclude  $H_a$  is true.

■ Greater than .10

Insufficient evidence to conclude  $H_a$  is true.

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#### **Describing the p-Value**

#### Example - Department Store's New Billing System

But how small does the p-value have to be to infer that the alternative hypothesis is true? The answer depends on several factors, including the costs of making Type I and Type II errors.

In the example, a Type I error would occur if the manager adopts the new billing system when it is not cost-effective.

If such cost is high, then we want to reject  $H_0$  when the p-value is quite low by setting a very low significance level, such as  $\alpha=1\%$ .

In general, the following descriptive terms apply:

• If p-value ≤ .01, there is overwhelming evidence that H₁ is true. The test is highly significant.

- If .01 < p-value ≤ .05, there is strong evidence that H₁ is true. The test is significant.</li>
- If .05 < p-value  $\leq$  .10, there is weak evidence that  $H_1$  is true. The test is weakly significant.
- If p-value > .10, there is little to no evidence that H<sub>1</sub> is true. The test is not significant.

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#### The p-Value and Rejection Region Methods

We can make similar decisions using either method.

When using the rejection region method, the following rules apply:

- If the test statistic is within the rejection region, we reject  $H_o$ .
- If the test statistic is outside the rejection region, we do not reject  $H_{\it o}$ .

When using the *p-value* method, the following rules apply:

- If *p*-value  $\leq \alpha$ , we reject  $H_o$ .
- If *p*-value >  $\alpha$ , we do not reject  $H_o$ .

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### **One- and Two-Tail Tests: Identify**

Example - The Effects of Kiosks on Sales in Fast-Food Restaurants

The example before was called a **one-tail test** because the rejection region is located only in one tail of the sampling distribution. We now present an example of a **two-tail test**.

A restaurant franchise, concerned about possible changes in sales due to the installation of self-serve kiosks, takes a random sample of 100 customers who use the new machine at McDonald's.

It is known that before the machine's installation, the average customer at McDonald's spent \$6.03 with a standard deviation of \$0.91. The sample mean is found to be equal to \$5.91.

Do the data provide enough evidence at the 5% significance level of a change in the size of the transaction for an individual customer?

Question: Has the mean expenditure changed from \$6.03?

And we know that  $\mu$  = 6.03,  $\sigma$  = 0.91, and  $\bar{x}$  =5.91

We set the hypotheses as:  $H_0$ :  $\mu = 6.03$  $H_1$ :  $\mu \neq 6.03$ 

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#### One- and Two-Tail Tests: Compute and Interpret

Because we can reject the null hypothesis when the test statistic is large or when it is small, and the total area in the rejection region must be  $\alpha$ , we divide this probability by 2:

$$z < z_{\alpha/2}$$
 or  $z > z_{\alpha/2}$ 

Because  $\alpha$  = .05, then  $\alpha$  / 2 = .025, and  $z_{.025}$  = 1.96.

Thus, we can write the rejection region as: z < -1.96 or z > 1.96

The value of the test statistic is:  $z = \frac{\bar{x} - \mu}{\sigma/\sqrt{n}} = \frac{5.91 - 6.03}{0.91/\sqrt{100}} = -1.32$ 

Because -1.32 is not in the rejection region, we cannot reject  $H_0$ .

The *p*-value corresponds to the area of both tails:

p-value = P(z < -1.32) + P(z > 1.32)=  $.0934 + .0934 = .1868 > \alpha$ 

There is not enough evidence to conclude that the self-serve kiosks result in a change from the amounts incurred with a cashier.



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### When Do We Conduct One- and Two-Tail Tests?

We conduct a **two-tail test** whenever we want to know whether there is enough evidence to infer that the mean is *not equal to* the value stated in the null hypothesis:

$$H_o$$
:  $\mu = \mu_o$ 

$$\mathrm{H}_{\scriptscriptstyle 1}\!\!:\mu\neq\mu_o$$

There are two one-tail tests.

We conduct a right one-tail test whenever we want to know whether there is enough evidence to infer that the mean is  $greater\ than$  the quantity specified by the null hypothesis:

$$H_o$$
:  $\mu = \mu_o$ 

$$H_1$$
:  $\mu > \mu_o$ 

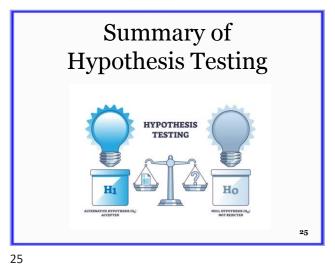
We conduct a  $\mathit{left}$  one-tail test whenever we want to know whether there is enough evidence to infer that the mean is  $\mathit{less}$  than the quantity specified by the null hypothesis:

$$H_o$$
:  $\mu = \mu_o$ 

$$H_1$$
:  $\mu < \mu_o$ 

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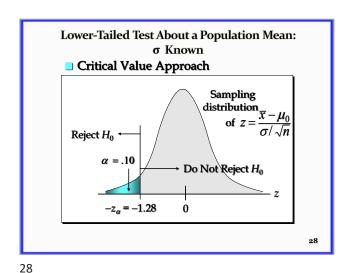


Critical Value Approach to **Two-Tailed Hypothesis Testing** 

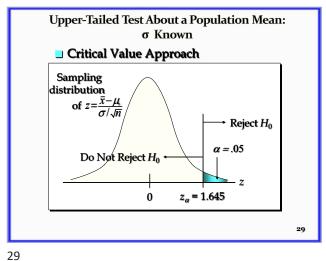
- The critical values will occur in both the lower and upper tails of the standard normal curve.
- Use the standard normal probability distribution table to find  $z_{\alpha/2}$  (the z-value with an area of  $\alpha/2$  in the upper tail of the distribution).
  - The rejection rule is: Reject  $H_0$  if  $z \le -z_{\alpha/2}$  or  $z \ge z_{\alpha/2}$ .

### Critical Value Approach to **One-Tailed Hypothesis Testing**

- The test statistic z has a standard normal probability distribution.
- We can use the standard normal probability distribution table to find the z-value with an area of  $\alpha$  in the lower (or upper) tail of the distribution.
- The value of the test statistic that established the boundary of the rejection region is called the critical value for the test.
  - The rejection rule is:
  - Lower tail: Reject  $H_0$  if  $z \leq -z_{\alpha}$
  - Upper tail: Reject H<sub>0</sub> if z ≥ z<sub>α</sub>



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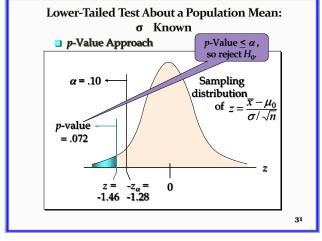
### p-Value Approach to **Two-Tailed Hypothesis Testing**

- Compute the *p*-value using the following three steps:
  - 1. Compute the value of the test statistic z.
  - 2. If z is in the upper tail (z > 0), find the area under the standard normal curve to the right of z.

If z is in the lower tail (z < 0), find the area under the standard normal curve to the left of z.

- 3. Double the tail area obtained in step 2 to obtain the p -value.
  - The rejection rule: Reject  $H_0$  if the p-value  $\leq \alpha$ .

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**Upper-Tailed Test About a Population Mean:** σ Known p-Value Approach p-Value  $\leq \alpha$ , so reject  $H_0$ . Sampling distribution  $\alpha = .04$ p-Value = .0110  $z_{\alpha} = 1.75$ 2.29 32

#### Summary of Forms for Null and Alternative Hypotheses about a Population Mean

- The equality part of the hypotheses always appears in the null hypothesis.
- In general, a hypothesis test about the value of a population mean  $\mu$  must take one of the following three forms (where  $\mu_0$  is the hypothesized value of the population mean).

$$H_0: \mu \ge \mu_0$$
  
 $H_a: \mu < \mu_0$ 

 $H_0$ :  $\mu \le \mu_0$  $H_a$ :  $\mu > \mu_0$ 

$$H_0: \mu = \mu_0$$

$$H_a: \mu \neq \mu_0$$

One-tailed (lower-tail)

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One-tailed (upper-tail)

Two-taile

33

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#### Confidence Interval Approach to Two-Tailed Tests About a Population Mean

- Select a simple random sample from the population and use the value of the sample mean  $\overline{x}$  to develop the confidence interval for the population mean  $\mu$ .
- If the confidence interval contains the hypothesized value μ<sub>0</sub>, do not reject H<sub>0</sub>. Otherwise, reject H<sub>0</sub>. (Actually, H<sub>0</sub> should be rejected if μ<sub>0</sub> happens to be equal to one of the end points of the confidence interval.)

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### **Testing Hypotheses and Confidence Interval**

Example - The Effects of Kiosks on Sales in Fast-Food Restaurants

The test statistic and the confidence interval estimator are both derived from the sampling distribution. It follows that we can use the confidence interval estimator to test hypotheses.

The 95% confidence interval estimate of the population mean is:

$$\bar{x} \pm z_{\alpha/2} \frac{\sigma}{\sqrt{n}} = 5.91 \pm 1.96 \frac{0.91}{\sqrt{100}} = 5.91 \pm 0.18$$
  
LCL = 5.73 and UCL = 6.09

We estimate that  $\mu$  lies between \$5.73 and \$6.09. Because the interval includes \$6.03, we cannot conclude that there is sufficient evidence to infer that the mean has changed.

This process is equivalent to the rejection region approach.

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**Steps of Hypothesis Testing Once Again** 

Step 1. Develop the null and alternative hypotheses.

**Step 2.** Specify the level of significance  $\alpha$ .

Step 3. Collect the sample data and compute the test statistic.

p-Value Approach

**Step 4.** Use the value of the test statistic to compute the *p*-value.

Step 5. Reject  $H_0$  if p-value  $\leq \alpha$ .

Critical Value Approach

**Step 4.** Use the level of significance to determine the critical value and the rejection rule.

Step 5. Use the value of the test statistic and the rejection rule to determine whether to reject  $H_0$ .

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### **Null and Alternative Hypotheses Example**

#### Metro EMS

A major west coast city provides one of the most comprehensive emergency medical services in the world. Operating in a multiple hospital system with approximately 20 mobile medical units, the service goal is to respond to medical emergencies with a mean time of 12 minutes or less.

The director of medical services wants to formulate a hypothesis test that could use a sample of emergency response times to determine whether or not the service goal of 12 minutes or less is being achieved.

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#### **Null and Alternative Hypotheses**

 $H_0: \mu \le 12$ 

The emergency service is meeting the response goal; no follow-up action is necessary.

 $H_a$ :  $\mu > 12$ 

The emergency service is not meeting the response goal; appropriate follow-up action is necessary.

where:  $\mu$  = mean response time for the population of medical emergency requests

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#### One-Tailed Tests About a Population Mean: σ Known

#### Metro EMS

The response times for a random sample of 40 medical emergencies were tabulated. The sample mean is 13.25 minutes. The population standard deviation is believed to be 3.2 minutes.

The EMS director wants to perform a hypothesis test, with a .05 level of significance, to determine whether the service goal of 12 minutes or less is being achieved.

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# One-Tailed Tests About a Population Mean: $\sigma$ Known

■ p-Value and Critical Value Approaches

1. Develop the hypotheses.

 $H_0$ :  $\mu \le 12$  $H_a$ :  $\mu > 12$ 

2. Specify the level of significance.

 $\alpha = .05$ 

3. Compute the value of the test statistic.

$$z = \frac{\overline{x} - \mu}{\sigma / \sqrt{n}} = \frac{13.25 - 12}{3.2 / \sqrt{40}} = 2.47$$

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# One-Tailed Tests About a Population Mean: $\sigma$ Known

■ p -Value Approach

#### 4. Compute the p -value.

For z = 2.47, cumulative probability = .9932. p-value = 1 - .9932 = (.0068)

#### 5. Determine whether to reject $H_0$ .

Because *p*-value = .0068  $\leq \alpha$  = .05, we reject  $H_0$ .

There is sufficient statistical evidence to infer that Metro EMS is <u>not</u> meeting the response goal of 12 minutes.

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# One-Tailed Tests About a Population Mean: $\sigma$ Known

- Critical Value Approach
  - 4. Determine the critical value and rejection rule.

For  $\alpha = .05$ ,  $z_{.05} = 1.645$ Reject  $H_0$  if  $z \ge 1.645$ 

5. Determine whether to reject  $H_0$ .

Because 2.47  $\geq$  1.645, we reject  $H_0$ .

There is sufficient statistical evidence to infer that Metro EMS is <u>not</u> meeting the response goal of 12 minutes.

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# Two-Tailed Tests About a Population Mean: $\sigma$ Known

**One-Tailed Tests About a Population Mean:** 

σKnown

0

1.645

2.47

 $\alpha = .05$ 

p-value

=.0068

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p -Value Approach

Sampling

distribution

**■ Example:** Glow Toothpaste

The production line for Glow toothpaste is designed to fill tubes with a mean weight of 6 oz. Periodically, a sample of 30 tubes will be selected in order to check the filling process.

Quality assurance procedures call for the continuation of the filling process if the sample results are consistent with the assumption that the mean filling weight for the population of toothpaste tubes is 6 oz.; otherwise the process will be adjusted.

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# Two-Tailed Tests About a Population Mean: $\sigma$ Known

Example: Glow Toothpaste

Assume that a sample of 30 toothpaste tubes provides a sample mean of  $6.1~\rm oz$ . The population standard deviation is believed to be  $0.2~\rm oz$ .

Perform a hypothesis test, at the .03 level of significance, to help determine whether the filling process should continue operating or be stopped and corrected.

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# Two-Tailed Tests About a Population Mean: $\sigma$ Known

- *p* -Value and Critical Value Approaches
  - Determine the hypotheses.

 $H_0$ :  $\mu = 6$  $H_a$ :  $\mu \neq 6$ 

2. Specify the level of significance.

 $\alpha = .03$ 

3. Compute the value of the test statistic.

$$z = \frac{\overline{x} - \mu_0}{\sigma / \sqrt{n}} = \frac{6.1 - 6}{.2 / \sqrt{30}} = 2.74$$

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45 46

# Two-Tailed Tests About a Population Mean: $\sigma$ Known

- p -Value Approach
  - 4. Compute the p -value.

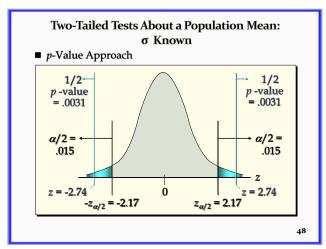
For z = 2.74, cumulative probability = .9969 p-value = 2(1 - .9969) = .0062

5. Determine whether to reject  $H_0$ .

Because *p*-value =  $.0062 \le \alpha = .03$ , we reject  $H_0$ .

There is sufficient statistical evidence to infer that the alternative hypothesis is true (i.e. the mean filling weight is not 6 ounces).

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# Two-Tailed Tests About a Population Mean: $\sigma$ Known

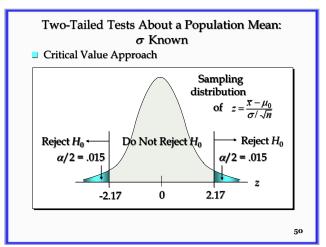
- Critical Value Approach
  - 4. Determine the critical value and rejection rule.

For 
$$\alpha/2 = .03/2 = .015$$
,  $z_{.015} = 2.17$   
Reject  $H_0$  if  $z \le -2.17$  or  $z \ge 2.17$ 

5. Determine whether to reject  $H_0$ .

Because 2.74  $\geq$  2.17, we reject  $H_0$ . There is sufficient statistical evidence to infer that the alternative hypothesis is true (i.e. the mean filling weight is not 6 ounces).

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### Confidence Interval Approach to Two-Tailed Tests About a Population Mean

The 97% confidence interval for  $\mu$  is

$$\overline{x} \pm z_{\alpha/2} \frac{\sigma}{\sqrt{n}} = 6.1 \pm 2.17 (.2/\sqrt{30}) = 6.1 \pm .07924$$

or 6.02076 to 6.17924

Because the hypothesized value for the population mean,  $\mu_0 = 6$ , is not in this interval, the hypothesis-testing conclusion is that the null hypothesis,  $H_0$ :  $\mu = 6$ , can be rejected.

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### A Summary of Forms for Null and Alternative Hypotheses About a Population Proportion

- The equality part of the hypotheses always appears in the null hypothesis.
- In general, a hypothesis test about the value of a population proportion *p* must take one of the following three forms (where ρ<sub>0</sub> is the hypothesized value of the population proportion).

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#### **Tests About a Population Proportion**

■ Test Statistic

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$$z = \frac{\bar{p} - p_0}{\sigma_{\bar{p}}}$$

where:

$$\sigma_{\bar{p}} = \sqrt{\frac{p_0(1-p_0)}{n}}$$

assuming  $np \ge 5$  and  $n(1-p) \ge 5$ 

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### **Tests About a Population Proportion**

■ Rejection Rule: p –Value Approach
Reject  $H_0$  if p –value  $\leq \alpha$ 

■ Rejection Rule: Critical Value Approach

$$H_0: p \le p_0$$
 Reject  $H_0$  if  $z \ge z_\alpha$ 

 $H_0$ :  $p \ge p_0$  Reject  $H_0$  if  $z \le -z_\alpha$ 

 $H_0$ :  $p = p_0$  Reject  $H_0$  if  $z \le -z_{\alpha/2}$  or  $z \ge z_{\alpha/2}$ 

### Two-Tailed Test About a Population Proportion

Example: National Safety Council (NSC)
 For a Christmas and New Year's week, the

National Safety Council estimated that 500 people would be killed and 25,000 injured on the nation's roads. The NSC claimed that 50% of the accidents would be caused by drunk driving.

A sample of 120 accidents showed that 67 were caused by drunk driving. Use these data to test the NSC's claim with  $\alpha$  = .05.

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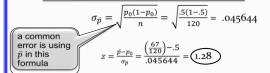
### Two-Tailed Test About a Population Proportion

■ p –Value and Critical Value Approaches

1. Determine the hypotheses.  $H_0: p = .5$  and  $H_a: p \neq .5$ 

2. Specify the level of significance.  $\alpha$  = .05

3. Compute the value of the test statistic.



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#### **Two-Tailed Test About a Population Proportion**

■ p-Value Approach

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4. Compute the *p* -value.

For z = 1.28, cumulative probability = .8997 p-value = 2(1 - .8997) = .2006

5. Determine whether to reject  $H_0$ .

Because *p*-value = .2006 >  $\alpha$  = .05, we cannot reject  $H_0$ .

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Interval Estimate of a Population Proportion

The general form of an interval estimate of a population proportion is

 $\overline{p} \pm \text{Margin of Error}$ 

The sampling distribution of  $\overline{P}$  plays a key role in computing the margin of error for this interval estimate.

The sampling distribution of  $\bar{P}$  can be approximated by a normal distribution whenever  $np \ge 5$  and  $n(1-p) \ge 5$ .

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Interval Estimate of a Population Proportion

Normal Approximation of Sampling Distribution of  $\overline{P}$ Sampling distribution of  $\overline{p}$   $\sigma_{\overline{p}} = \sqrt{\frac{p(1-p)}{n}}$ 

Interval Estimate of a Population Proportion

**■** Interval Estimate

 $\overline{p} \pm z_{\alpha/2} \sqrt{\frac{\overline{p}(1-\overline{p})}{n}}$ 

where:  $1-\alpha$  is the confidence coefficient  $z_{\alpha/2}$  is the z value providing an area of  $\alpha/2$  in the upper tail of the standard normal probability distribution is the sample proportion

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#### **Interval Estimate of a Population Proportion**

**Example:** Survey of women golfers

A national survey of 900 women golfers was conducted to learn how women golfers view their treatment at golf courses in United States. The survey found that 396 of the women golfers were satisfied with the availability of tee times.

Suppose one wants to develop a 95% confidence interval estimate for the proportion of the population of women golfers satisfied with the availability of tee times.

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#### **Interval Estimate of a Population Proportion**

**Example:** Survey of women golfers

$$\bar{p} \pm z_{\alpha/2} \sqrt{\frac{\bar{p}(1-\bar{p})}{n}}$$

where: n = 900,  $\bar{p} = 396/900 = .44$ ,  $z_{\alpha/2} = 1.96$ 

$$.44 \pm 1.96 \sqrt{\frac{.44(1-.44)}{900}} = .44 \pm .0324$$

Survey results enable us to state with 95% confidence that between 40.76% and 47.24% of all women golfers are satisfied with the availability of tee times.

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### Sample Size for an Interval Estimate of a Population Proportion

Margin of Error

$$E = z_{\alpha/2} \sqrt{\frac{\bar{p}(1-\bar{p})}{n}}$$

Solving for the necessary sample size n, we get

$$n = \frac{\left(z_{\alpha/2}\right)^2 \bar{p}(1-\bar{p})}{F^2}$$

However,  $\bar{p}$  will not be known until after we have selected the sample. We will use the planning value  $p^*$  for  $\bar{p}$ .

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# Sample Size for an Interval Estimate of a Population Proportion

Necessary Sample Size

$$n = \frac{\left(z_{\alpha/2}\right)^2 p^* (1 - p^*)}{F^2}$$

The planning value  $p^*$  can be chosen by:

- Using the sample proportion from a previous sample of the same or similar units, or
- Selecting a preliminary sample and using the sample proportion from this sample.
- 3. Using judgment or a "best guess" for a  $p^*$  value.
- 4. Otherwise, using .50 as the  $p^*$  value.

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# Sample Size for an Interval Estimate of a Population Proportion

**Example:** Survey of women golfers

Suppose the survey director wants to estimate the population proportion with a margin of error of .025 at 95% confidence.

How large a sample size is needed to meet the required precision? (A previous sample of similar units yielded .44 for the sample proportion.)

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Sample Size for an Interval Estimate of a Population Proportion

**Example:** Survey of women golfers

$$E = z_{\alpha/2} \sqrt{\frac{p^*(1-p^*)}{n}} = .025$$

At 95% confidence,  $z_{.0125} = 1.96$ . Recall that  $p^* = .44$ .

$$n = \frac{\left(z_{\alpha/2}\right)^2 p^* (1 - p^*)}{E^2} = \frac{(1.96)^2 (.44)(.56)}{(.025)^2} = 1514.5$$

A sample of size 1515 is needed to reach a desired precision of ± .025 at 95% confidence.

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# Interval Estimate of a Population Proportion (another example)

Example: Political Science, Inc.

Political Science, Inc. (PSI) specializes in voter polls and surveys designed to keep political office seekers informed of their position in a race. Using telephone surveys, PSI interviewers ask registered voters who they would vote for if the election were held that day.

In a current election campaign, PSI has just found that 200 registered voters, out of 500 contacted, favor a particular candidate. PSI wants to develop a 90% confidence interval estimate for the proportion of the population of registered voters that favor the candidate.

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Interval Estimate of a Population Proportion  $\overline{p(1-p)}$ 

$$\overline{p} \pm z_{\alpha/2} \sqrt{\frac{\overline{p}(1-\overline{p})}{n}}$$

where: n = 500,  $\overline{p} = 200/500 = .40$ ,  $z_{\alpha/2} = 1.645$ 

$$.40 \pm 1.645 \sqrt{\frac{.40(1 - .40)}{500}} = .40 \pm .036$$

PSI is 90% confident that the proportion of all voters that favor the candidate is between .364 and .436.

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#### Sample Size for an Interval Estimate of a Population Proportion (another example)

Example: Political Science, Inc.

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Suppose that PSI would like a .99 probability that the sample proportion is within + .03 of the population proportion.

How large a sample size is needed to meet the required precision? (A previous sample of similar units yielded .40 for the sample proportion.)

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Sample Size for an Interval Estimate of a Population Proportion

$$z_{\alpha/2}\sqrt{\frac{p(1-p)}{n}}=.03$$
 At 99% confidence,  $z_{.005}$  = 2.576. Recall that  $\bar{p}$  = .40.

$$n = \frac{(z_{\alpha/2})^2 p(1-p)}{E^2} = \frac{(2.576)^2 (.40)(.60)}{(.03)^2} \cong \boxed{1769}$$

A sample of size 1769 is needed to reach a desired precision of ± .03 at 99% confidence.

#### **Stata Output and Instructions** Example 11.1 - Testing a Population Mean: Standard Deviation Known Instructions 1. Import the data into one column. (Click File/Import/Excel spreadsheet(\*xls,\*xlsx)/Chapter11/Xm11-01.) Check Import first row as variable names. Click Statistics, Summaries, tables and tests, Classical tests of hypotheses, and z-test (mean comparison test known variance). 3. Select One-sample, select Accounts in the Variable name: box and type 170 in the Hypothesized mean: box. Type 65 for the Standard deviation. Variable Mean Std. Err. Std. Dev. [95% Conf. Interval] mean = mean(Accounts) Ho: mean = 170 z = 2,4605 Ha: mean != 170 Pr(|Z| > |z|) = 0.0139 Ha: mean < 170 Pr(Z < z) = 0.9931Pr(Z > z) = 0.0069

**Stata Output and Instructions** Example 11.2 – Testing a Population Mean: Standard Deviation Known Instructions: 1. Import the data into one column. (Click File/Import/Excel spreadsheet(\*xls,\*xlsx)/Chapter11/Xm11-02.) Check Import first row as variable Click Statistics, Summaries, tables and tests, Classical tests of hypotheses, and z-test (mean comparison test known variance). 3. Select One-sample, select Accounts in the Variable name: box and type 6.03 in the Hypothesized mean: box. Type 0.91 for the Standard deviation Variable Mean Std. Err. Std. Dev. [95% Conf. Interval] Keller, Gerald, Statistics for Management and Economics, 12th Edition. © 2023 Cengage. All Rights Reserved